

L7: Kernel density estimation

Non-parametric density estimation

Histograms

Parzen windows

Smooth kernels

Product kernel density estimation

The naïve Bayes classifier

Density Functions

Suppose we have some variable $X \sim f(x)$ where $f(x)$ is the **probability density function (pdf)** of X .

Note that we have two requirements on $f(x)$:

- $f(x) \geq 0$ for all $x \in \mathcal{X}$, where \mathcal{X} is the domain of X
- $\int_{\mathcal{X}} f(x)dx = 1$

Example: normal distribution pdf has the form

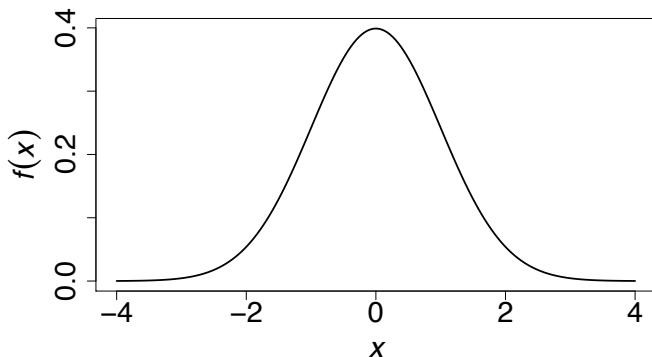
$$f(x) = \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$$

which is well-defined for all $x, \mu \in \mathbb{R}$ and $\sigma \in \mathbb{R}^+$.

Standard Normal Distribution

If $X \sim N(0, 1)$, then X follows a standard normal distribution:

$$f(x) = \frac{1}{\sqrt{2\pi}} e^{-x^2/2} \quad (1)$$



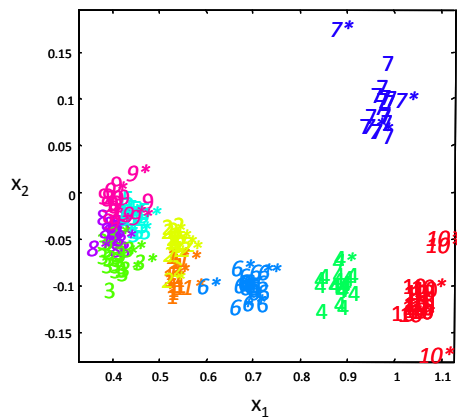
Non-parametric density estimation

In the previous two lectures we have assumed that either

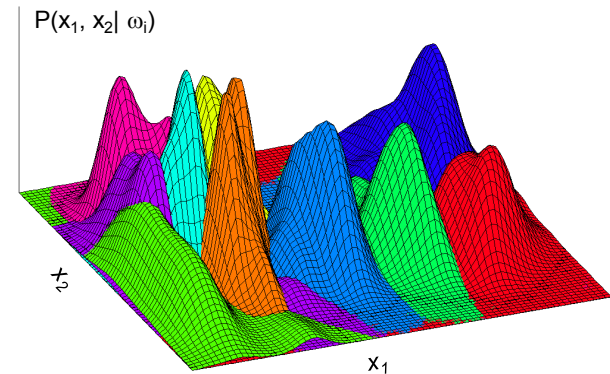
- The likelihoods $p(x|\omega_i)$ were known (LRT), or
- At least their parametric form was known (parameter estimation)

The methods that will be presented in the next two lectures do not afford such luxuries

- Instead, they attempt to estimate the density directly from the data without assuming a particular form for the underlying distribution
- Sounds challenging? You bet!



→
non-parametric
density estimation



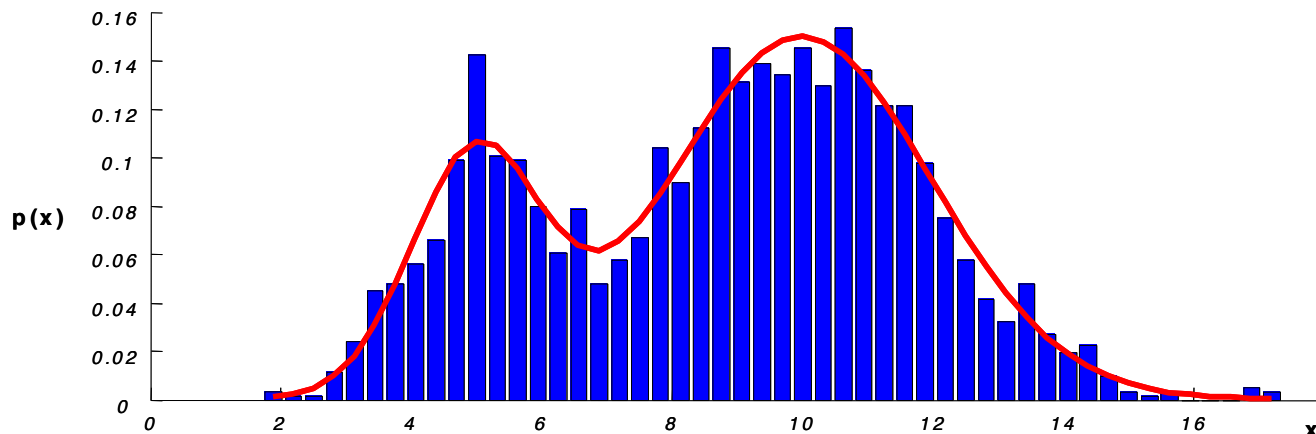
The histogram

The simplest form of non-parametric DE is the histogram

- Divide the sample space into a number of bins and approximate the density at the center of each bin by the fraction of points in the training data that fall into the corresponding bin

$$p_H(x) = \frac{1}{N} \frac{[\# \text{ of } x^{(k)} \text{ in same bin as } x]}{[\text{width of bin}]}$$

- The histogram requires two “parameters” to be defined: bin width and starting position of the first bin



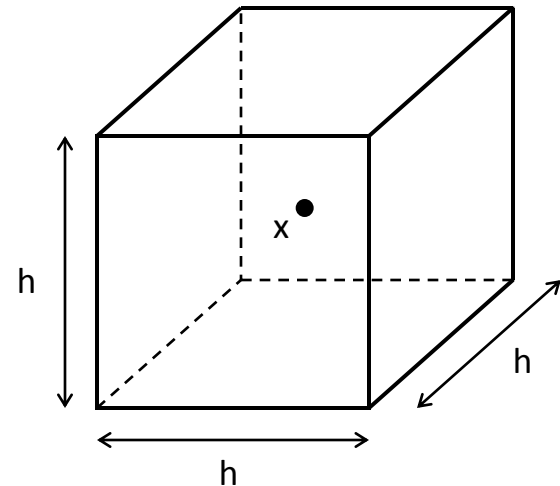
The histogram is a very simple form of density estimation, but has several drawbacks

- The density estimate depends on the starting position of the bins
 - For multivariate data, the density estimate is also affected by the orientation of the bins
- The discontinuities of the estimate are not due to the underlying density; they are only an artifact of the chosen bin locations
 - These discontinuities make it very difficult (to the naïve analyst) to grasp the structure of the data
- A much more serious problem is the curse of dimensionality, since the number of bins grows exponentially with the number of dimensions
 - In high dimensions we would require a very large number of examples or else most of the bins would be empty
- These issues make the histogram unsuitable for most practical applications except for quick visualizations in one or two dimensions
- Therefore, we will not spend more time looking at the histogram

Parzen windows

Problem formulation

- Assume that the region \mathfrak{R} that encloses the k examples is a hypercube with sides of length h centered at x
 - Then its volume is given by $V = h^D$, where D is the number of dimensions



- To find the number of examples that fall within this region we define a kernel function $K(u)$

$$K(u) = \begin{cases} 1 & |u_j| < 1/2 \quad \forall j = 1 \dots D \\ 0 & \text{otherwise} \end{cases}$$

- This kernel, which corresponds to a unit hypercube centered at the origin, is known as a Parzen window or the naïve estimator
- The quantity $K((x - x^{(n)})/h)$ is then equal to unity if $x^{(n)}$ is inside a hypercube of side h centered on x , and zero otherwise

[Bishop, 1995]

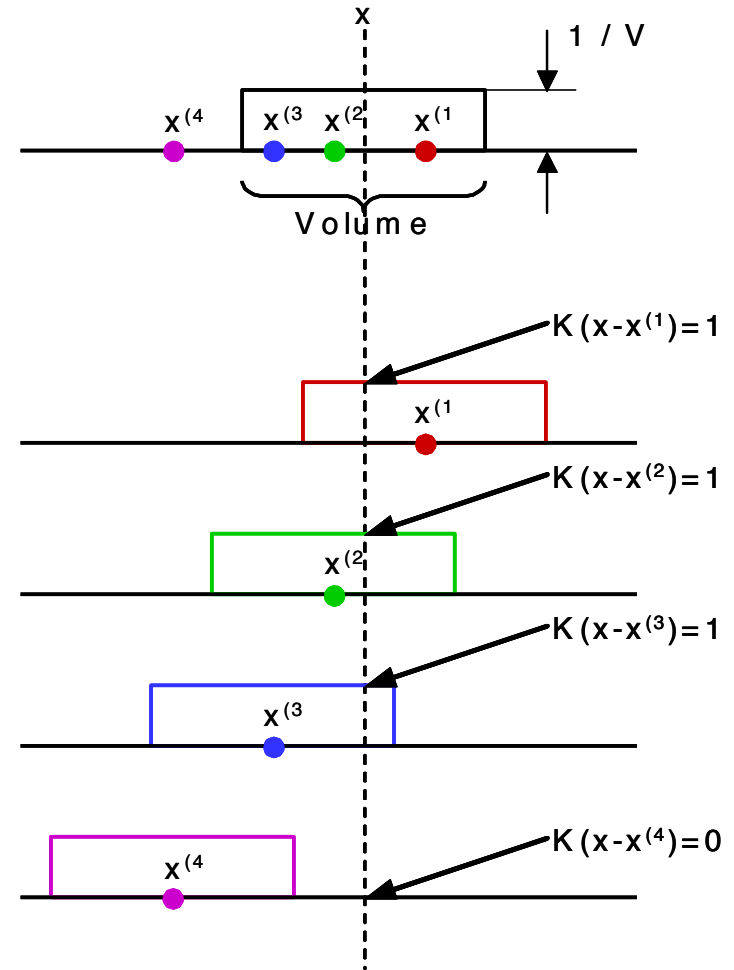
- The total number of points inside the hypercube is then

$$k = \sum_{n=1}^N K\left(\frac{x - x^{(n)}}{h}\right)$$

Substituting back into the expression for the density estimate

$$p_{KDE}(x) = \frac{1}{Nh^D} \sum_{n=1}^N K\left(\frac{x - x^{(n)}}{h}\right)$$

- Notice how the Parzen window estimate resembles the histogram, with the exception that the bin locations are determined by the data



Smooth kernels

The Parzen window has several drawbacks

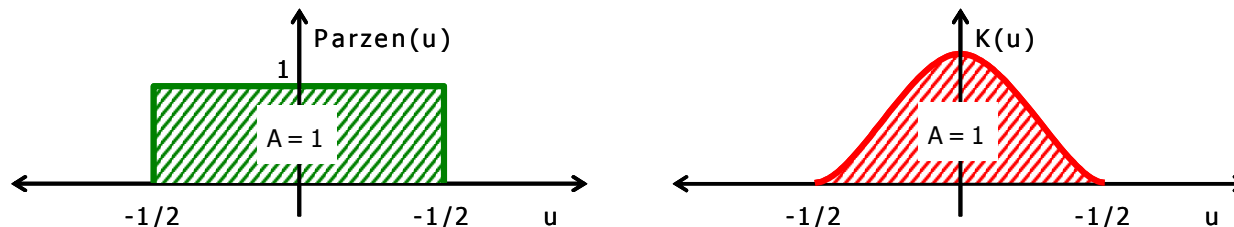
- It yields density estimates that have discontinuities
- It weights equally all points x_i , regardless of their distance to the estimation point x

For these reasons, the Parzen window is commonly replaced with a smooth kernel function $K(u)$

$$\int_{\mathbb{R}^D} K(x) dx = 1$$

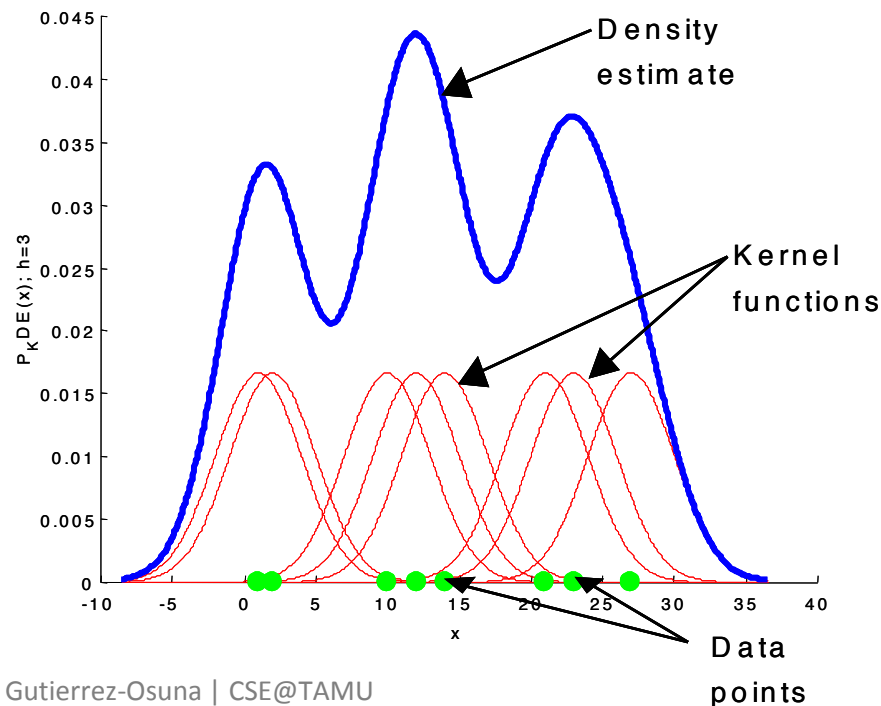
- Usually, but not always, $K(u)$ will be a radially symmetric and unimodal pdf, such as the Gaussian $K(x) = (2\pi)^{-D/2} e^{-\frac{1}{2}x^T x}$
- Which leads to the density estimate

$$p_{KDE}(x) = \frac{1}{Nh^D} \sum_{n=1}^N K\left(\frac{x-x^{(k)}}{h}\right)$$



Interpretation

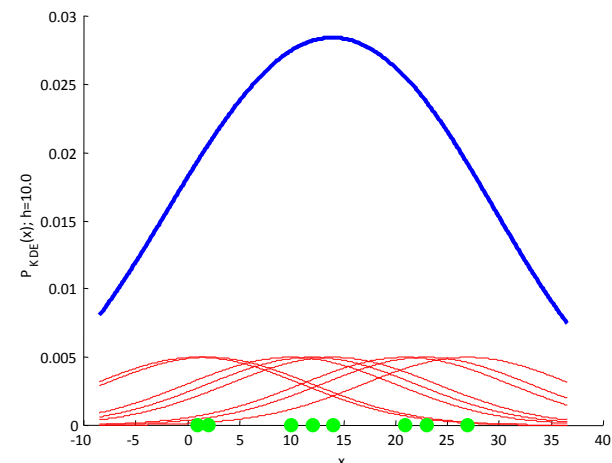
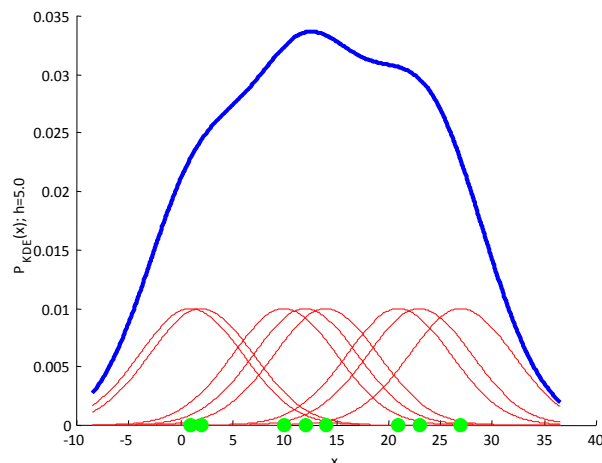
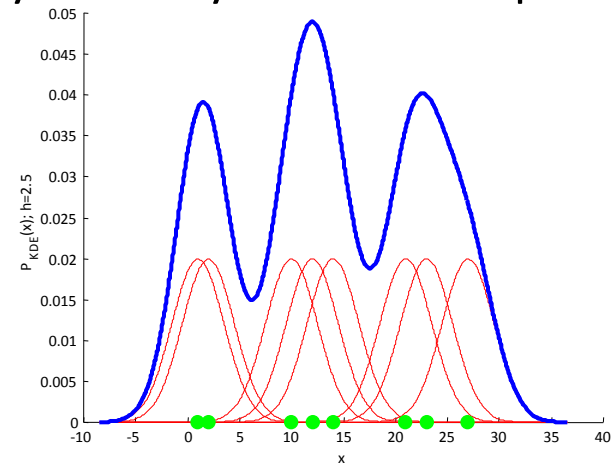
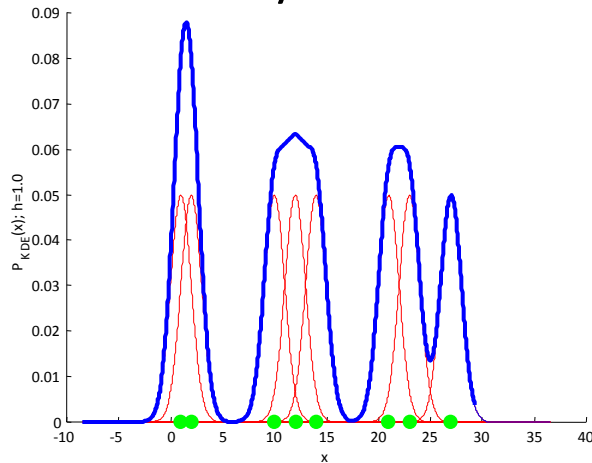
- Just as the Parzen window estimate can be seen as a sum of boxes centered at the data, the smooth kernel estimate is a sum of “bumps”
- The kernel function determines the shape of the bumps
- The parameter h , also called the smoothing parameter or bandwidth, determines their width



Bandwidth selection

The problem of choosing h is crucial in density estimation

- A large h will over-smooth the DE and mask the structure of the data
- A small h will yield a DE that is spiky and very hard to interpret



Multivariate density estimation

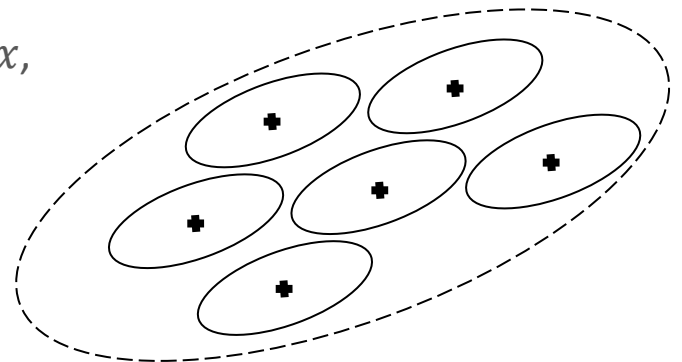
For the multivariate case, the KDE is

$$p_{KDE}(x) = \frac{1}{Nh^D} \sum_{n=1}^N K\left(\frac{x-x^{(n)}}{h}\right)$$

- Notice that the bandwidth h is the same for all the axes, so this density estimate will be weight all the axis equally
- If one or several of the features has larger spread than the others, we should use a vector of smoothing parameters or even a full covariance matrix, which complicates the procedure

There are two basic alternatives to solve the scaling problem without having to use a more general KDE

- Pre-scaling each axis (normalize to unit variance, for instance)
- Pre-whitening the data (linearly transform so $\Sigma = I$), estimate the density, and then transform back [Fukunaga]
 - The whitening transform is $y = \Lambda^{-1/2} M^T x$, where Λ and M are the eigenvalue and eigenvector matrices of Σ
 - Fukunaga's method is equivalent to using a hyper-ellipsoidal kernel



Product kernels

A good alternative for multivariate KDE is the product kernel

$$p_{PKDE}(x) = \frac{1}{N} \sum_{i=1}^N K(x, x^{(n)}, h_1, \dots, h_D)$$

$$\text{where } K(x, x^{(n)}, h_1, \dots, h_D) = \frac{1}{h_1 \dots h_D} \prod_{d=1}^D K_d \left(\frac{x_d - x_d^{(n)}}{h_d} \right)$$

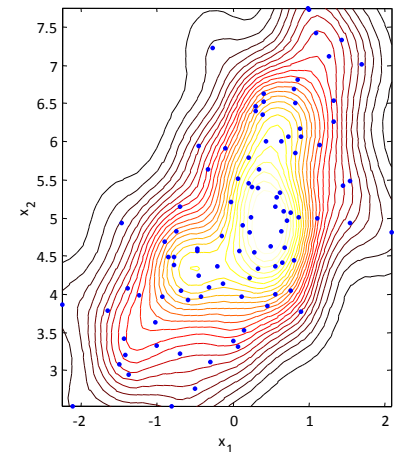
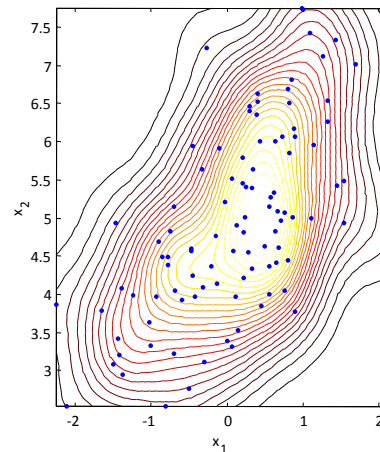
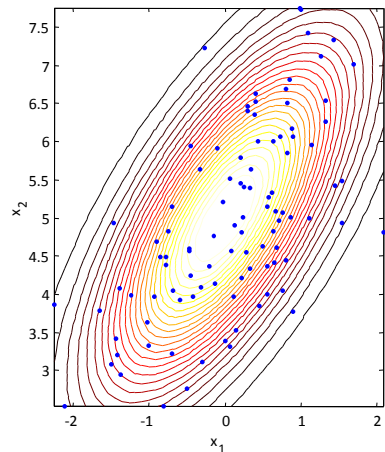
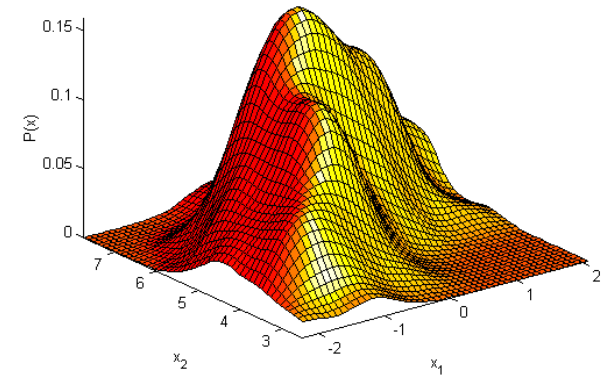
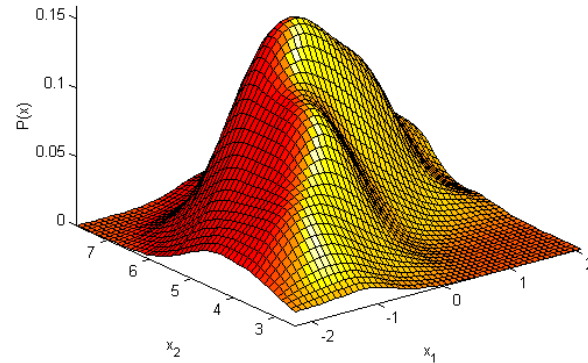
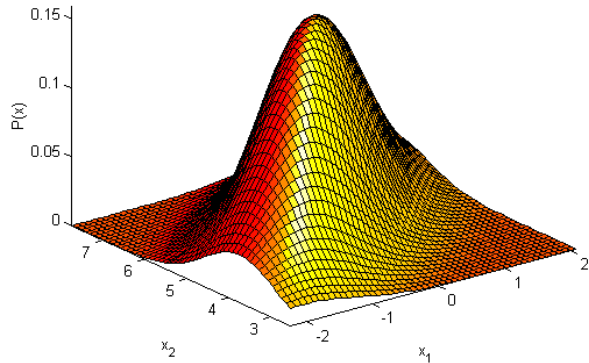
- The product kernel consists of the product of one-dimensional kernels
 - Typically the same kernel function is used in each dimension ($K_d(x) = K(x)$), and only the bandwidths are allowed to differ
 - Bandwidth selection can then be performed with any of the methods presented for univariate density estimation
- Note that although $K(x, x^{(n)}, h_1, \dots, h_D)$ uses kernel independence, this does not imply we assume the features are independent
 - If we assumed feature independence, the DE would have the expression

$$p_{FEAT-IND}(x) = \prod_{d=1}^D \frac{1}{Nh^D} \sum_{i=1}^N K_d \left(\frac{x_d - x_d^{(n)}}{h_d} \right)$$

- Notice how the order of the summation and product are reversed compared to the product kernel

Example I

- This example shows the product KDE of a bivariate unimodal Gaussian
 - 100 data points were drawn from the distribution
 - The figures show the true density (left) and the estimates using $h = 1.06\sigma N^{-1/5}$ (middle) and $h = 0.9AN^{-1/5}$ (right)



Example II

- This example shows the product KDE of a bivariate bimodal Gaussian
 - 100 data points were drawn from the distribution
 - The figures show the true density (left) and the estimates using $h = 1.06\sigma N^{-1/5}$ (middle) and $h = 0.9AN^{-1/5}$ (right)

